

Maryland Longitudinal Data System

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School Diversity, Academic Outcomes, and the Operationalization of Race-Ethnicity Peer Effects

Francisco Lagos
MLDS Research Branch
UMCP



Motivation

- Recent demographic changes in the US → increased diversity in the school system
- Peer effects literature has divergent perspectives on the effects of school diversity, especially for students of color
 - Some argue for potential benefits of diversity associated with access to resources found predominantly in white or affluent schools (Harris et al., 2006; Herring, 2009)
 - Others propose that homogeneous schooling environments may better target population's needs (Braddock & Eitle, 2004; Mickelson & Nkomo, 2012)

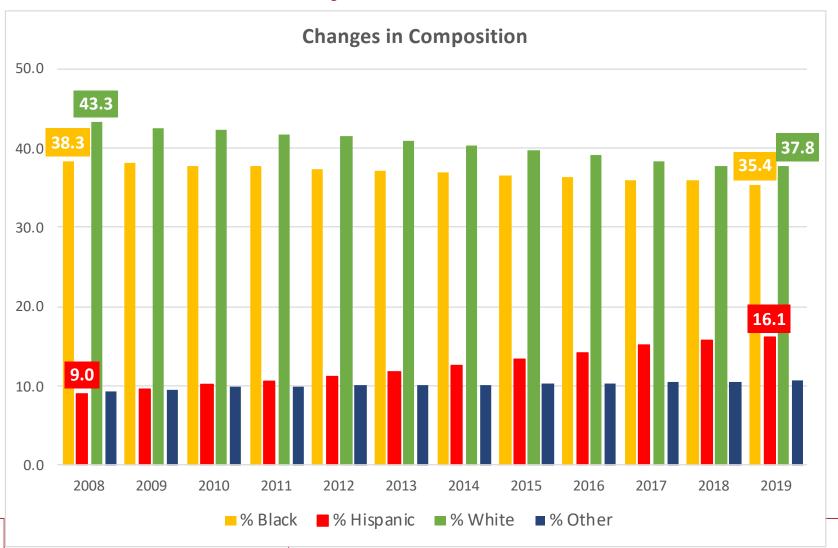


Research Questions

- Gaps in the literature
 - Operationalization of peer effects
 - Long term effects
- Research questions
 - What is the relationship between racial/ethnic diversity in schools and academic outcomes? How does this relationship change depending on the operationalization of peer characteristics and diversity?



Context: Maryland





Data and Measures from MLDS

- 2007-08 through 2018-19 school years
- Outcomes:
 - Short-term: math and ELA test scores, ever suspended
 - Long-term: college enrollment (at age 20), college graduation (2-yr or 4-yr program, at age 23)
- Independent variables:
 - School percentage of students of different race-ethnicities
 - o 3 models of peer effects:
 - Linear-in-means: % students of color (continuous)
 - Boutique/tracking: majority of students of the same raceethnicity (dichotomic)
 - Rainbow: different populations represented in schools (dichotomic)



Empirical Strategy and Sample

- Short-term effects:
 - Student, school, and year fixed effect (robustness: similar results when adding grade fixed effects; potential extension: value added approach)
 - Sample: elementary students (extension/robustness: running results with all K-12 students, but so far similar)
- Long-term effects:
 - School and year fixed effects
 - Sample: 9th grade students (extension: working on starting with other grade level high school cohorts)
- Student (gender, ELL, special education, and FARMs status) and school control variables (total enrollment, proportion FARMs)
- Results clustered at the school level



Limitations

- Long term models cannot include student fixed effects (no variation in the outcome at the individual level)
 - Potential bias due to endogenous assignment of students to peers.
 - I partially deal with this by including lagged test scores and suspensions as control variables.
- Long term estimations focus on college enrollment and graduation.
 - Due to sample restrictions in the wages data (only individuals working in Maryland and only in some sectors), I do not include wages in my estimations.



Preliminary Results

- I focus on the three largest groups of students in Maryland, i.e.,
 Black, Hispanic, and white.
- In the short term, students from these race-ethnicity groups seem to benefit from being with peers like themselves
 - Black students' math and ELA test scores increase as the % of students of color increase or when enrolled in majority Black schools.
 - Hispanic students' suspensions decrease as the % of students of color increase.
 - White students' ELA test scores increase when enrolled in majority white schools.
- Results are robust to different cut off points for boutique/tracking and rainbow models of peer effects



Short-Term

| | Black | | | Hispanic | | | White | | |
|-----------------------|----------|----------|-------------------|----------|----------|-------------------|----------|----------|-------------------|
| | Math | ELA | Ever suspended | Math | ELA | Ever suspended | Math | ELA | Ever suspended |
| Linear-in-means | | | | | | | , | | |
| % student of color | 0.0041** | 0.0027* | -0.0003 | 0.0028 | 0.0018 | -0.0003* | -0.0009 | -0.0019 | 0.0001 |
| | (0.0013) | (0.0011) | (0.0002) | (0.0019) | (0.0016) | (0.0001) | (0.0014) | (0.0011) | (0.0001) |
| Boutique/tracking | | | _ | | • | | • | | |
| Majority same | 0.0171 | 0.0303* | -0.0019 | -0.0375 | -0.0058 | 0.0013 | 0.0100 | 0.0300* | -0.0005 |
| race/ethnicity | (0.0214) | (0.0152) | (0.004) | (0.0266) | (0.0214) | (0.0016) | (0.0164) | (0.0142) | (0.0011) |
| Rainbow | • | | | | | | • | | |
| Black/Hispanic, white | -0.0171 | -0.0086 | 0.0018 | -0.0178 | -0.0204 | 0.0009 | -0.0109 | -0.0022 | -0.0003 |
| representative | (0.0139) | (0.0109) | (0.0032) | (0.0189) | (0.015) | (0.0013) | (0.0131) | (0.011) | (0.0012) |
| Observations | 594,640 | 592,940 | 1,271,435 | 265,540 | 260,221 | 600,989 | 701,943 | 700,340 | 1,489,867 |
| Covariates | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ |
| Fixed effects | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ |



Preliminary Results

- In the long term, the association between peer composition and academic outcomes goes in the opposite direction
 - Black and Hispanic students' college enrollment and/or graduation from college decreases as the proportion of peers of color in their 9th grade school cohort increases or when enrolled in majority same-race schools in 9th grade.
 - Hispanic and white students' college graduation increases when enrolled in diverse/representative schools in 9th grade.
- Again, results are robust to different cut off points for boutique/tracking and rainbow models of peer effects



Long-Term

| | Black | | | Hispanic | | | White | | |
|-----------------------|------------|-----------|-----------|------------|------------|-----------|------------|-----------|---------------|
| | Enrollment | Grad 2 yr | Grad 4 yr | Enrollment | Grad 2 yr | Grad 4 yr | Enrollment | Grad 2 yr | Grad 4 yr |
| | | program | program | | program | program | | program | program |
| Linear-in-means | | _ | | | | | | | |
| % student of color | -0.0028* | -0.0015 | -0.0039 | -0.0017 | 0.0016 | -0.0086** | 0.0006 | 0.0008 | -0.0032 |
| | (0.0011) | (0.0009) | (0.0025) | (0.0019) | (0.0025) | (0.0030) | (0.0010) | (0.0014) | (0.0026) |
| Boutique/tracking | | • | _ | | | | • | | |
| Majority same | -0.0012 | -0.0131* | 0.0160 | -0.0191 | -0.0837*** | -0.0036 | -0.0037 | -0.0005 | 0.0037 |
| race/ethnicity | (0.0104) | (0.0052) | (0.0125) | (0.0232) | (0.0063) | (0.0099) | (0.0111) | (0.0078) | (0.0079) |
| Rainbow | | | _ | | | | | | |
| Black/Hispanic, white | -0.0114 | -0.0309 | 0.0205 | 0.0118 | -0.0003 | 0.1194*** | 0.0024 | 0.0282*** | 0.0008 |
| representative | (0.0198) | (0.0288) | (0.0191) | (0.0310) | (0.0208) | (0.0285) | (0.0072) | (0.0049) | (0.0248) |
| Observations | 111,432 | 47,858 | 47,858 | 29,559 | 12,021 | 12,021 | 135,336 | 57,948 | <i>57,948</i> |
| Covariates | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ |
| Fixed effects | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ | Υ |



Discussion and Next Steps

- Operationalization of peer effects is relevant
 - O Linear-in-means models provide only a partial picture (e.g., if looking only at the linear association between white students and the % of peers of color, we may conclude that composition is not relevant in the short or long run)
- Mixed evidence between short- and long-term effects points to the need of understanding the mechanisms
 - For short-term effects, one mechanism to explore is the proportion of same-race teachers/staff → As the percentage of students of color increases, the likelihood of having a teacher of color also increases. In diverse (rainbow) schools, however, that likelihood decreases.
- Next steps: 1) complete the analysis, extensions, and robustness; 2) explore mechanisms.



Thank you!

Francisco Lagos: <u>flagosm@umd.edu</u>; @lagospancho